

A NON-LINEAR ORDER OF THE MARKOV PROCESS

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Abstract

The non-linear order $M < \infty$ of the Markov process $\{X(t), 0 \leq t\}$ is defined as $M = \sup M(s)$ where $M(s)$ is the minimal number of random variables generating $\sigma\{E_s X(t), s \leq t\}$ ($E_s(\cdot)$ is the conditional expectation with respect to $\sigma\{X(u), u \leq s\}$). The relation between the non-linear and linear order of the Markov process is discussed. The latter is a modification of the Hida multiple Markov process.

Introduction

According to [1] a Gaussian process $\{X(t), 0 \leq t\}$ is an N -ple Markov process if for any $0 < s \leq t_1 < \dots < t_N$, $E_s(X(t_i)), i=1, \dots, N$ are linearly independent and for any $0 < s \leq t_1 < \dots < t_{N+1}$, $E_s(X(t_i)), i=1, \dots, N+1$ are linearly dependent. $E_s(\cdot)$ is the conditional expectation with respect to the σ -field $\mathfrak{F}_s(x)$ generated by $\{X(u), u \leq s\}$. A small modification of this definition is given in [8]: Consider the mean-square linear closure \tilde{K}_b^a of $\{E_s(X(t)), a \leq b \leq t\}$. The Gaussian process $\{X(t), 0 \leq t\}$ is an N -ple Markov process if for any $0 < a \leq b$ the dimension of the space \tilde{K}_b^a is exactly N .

The following modification of this definition seems to be convenient: Let \tilde{K}_t be the mean-square linear closure of $\{E_s(X(t)), s \leq t\}$ and let $M(s)$ be the dimension of \tilde{K}_t . The Gaussian process $\{X(t), 0 \leq t\}$ is an N -ple Markov process if $N = \sup M(s)$ is finite.

For example, consider the process $\{\xi(t), 0 \leq t \leq 1\}$ defined by

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$$\xi(t) = W_1(t) + \varphi(t)W_2(t)$$

where $\varphi(t)$ is a version of the Cantor distribution function and $\{W_i(t), 0 \leq t \leq 1\}$, $i=1,2$ are independent Wiener processes. It is shown in [7] that $M(s)=2$ for s left-end points of removed intervals in the construction of the Cantor ternary set and $M(s)=3$ for the other s in $[0,1]$.

Let $\{X(t), 0 \leq t\}$ be a second-order process with a separable space $\mathcal{H}_t(x)$ - the mean-square linear closure of $\{X(u), u \leq t\}$ and let $\bigcap_t \mathcal{H}_t(x) = 0$. Let $P_s(\cdot)$ be the projection operator onto $\mathcal{H}_t(x)$. Let $M(s)$ be the dimension of the space \mathcal{H}_s - the mean-square closure of $\{P_s(X(t)), s \leq t\}$ and let $N = \sup M(s)$.

Definition 1. The process $\{X(t), 0 \leq t\}$ is the Markovian of the linear order N , if N is finite.

Now, let $\{X(t), 0 \leq t\}$ be the process with $E(X(t)) = 0$. Let \mathfrak{F}_s be the σ -field generated by $\{E_s(X(t)), s \leq t\}$. Denote by $M(s)$ the minimal number of random variables generating \mathfrak{F}_s and let $N = \sup M(s)$.

Definition 2. The process $\{X(t), 0 \leq t\}$ is the Markovian of the non-linear order N , if N is finite.

We find in [2] and [5] some motivations for Definition 2. In these papers the Gaussian Markov N -ple process $\{\xi(t), 0 \leq t\}$ was considered for which, $0 \leq s \leq t$,

$$E_s(\xi(t)) = \sum_{j=0}^{N-1} a_j(s,t) \xi^{(j)}(s)$$

and the process $\{X(t), 0 \leq t\}$ defined by $X(t) = \xi^n(t)$, $n \geq 2$. It was shown that for $n=2$ ([2]) $E_s(X(t))$ is the rational function of $X(s), X'(s), \dots, X^{(N-1)}(s)$, and for $n \geq 3$ ([5])

$$E_s(X(t)) = X(s)R_0 + \left[X'(s)\right]^{n-2} R_1 + \dots + \left[X^{(N)}(s)\right]^{n-4} R_2 + \dots,$$

where R_0, R_1, R_2, \dots is the rational function of $X(s), X'(s), \dots, X^{(N-1)}(s)$.

We shall discuss the relation between the linear and non-linear order by some examples subsequently in the paper. These examples are mainly based on the Hermite polynomials $H_p(\xi_1, \xi_2, \dots, \xi_p)$ of the Gaussian variables ξ_1, ξ_2, \dots . Especially, the property that the conditional expectation and arbitrary Hermite polynomial commute:

$$E_{\mathbf{a}}(H_p(\xi(t_1)), \dots, H_p(\xi(t_p))) = H_p(E_{\mathbf{a}}(\xi(t_1)), \dots, E_{\mathbf{a}}(\xi(t_p))). \quad (41).$$

Some properties of the Markov border and the multiplicity of the Hermite polynomials of the Gaussian Markov process and fields are treated in [6].

First we shall remark that Definitions 1 and 2 are mutually free from one another. Indeed, for a Markov process $\{X(t)\}$ of the non-linear order N not to be of some linear order M it is sufficient that $E(X^2(t)) = \infty$. Conversely, there are Markov processes of a linear order N , which are not Markov processes of some non-linear order M , as the following example shows.

Example 1. Let ξ_1, ξ_2, \dots be a sequence of Gaussian variables with a complete stochastic dependence and let $\{X(t), 0 \leq t\}$ be defined by

$$X(t) = \begin{cases} 0 & , \quad t < 1 \\ \sum_{p=1}^{[t]} H_p(\xi_p), & 1 \leq t \end{cases} \quad , \quad (H_p(\xi) = H_p(\underbrace{\xi, \dots, \xi}_p)) .$$

$\{X(t)\}$ is a process with orthogonal increments, so it is the Markovian of the linear order $M=1$. But, for $1 \leq s < t$

$$\begin{aligned} E_{\mathbf{a}}(X(t)) &= E \left[\sum_{p=1}^{[t]} H_p(\xi_p) \mid \xi_1, \dots, \xi_{[s]} \right] = \sum_{p=1}^{[s]} H_p(\xi_p) + \\ &+ E \left[\sum_{p=[s]+1}^{[t]} H_p(\xi_p) \mid \xi_1, \dots, \xi_{[s]} \right] = X(s) + \sum_{p=[s]+1}^{[t]} H_p(E(\xi_p \mid \xi_1, \dots, \xi_{[s]})) . \end{aligned}$$

Hence, $M(s) = [s]$ and $M = \infty$. \square

It is a well-known fact that $E_{\mathbf{a}} = P_{\mathbf{a}}$ for Gaussian processes. The next two examples show that the class of processes for which $E_{\mathbf{a}} = P_{\mathbf{a}}$ is large.

Example 2. Let $\{\xi(t), 0 \leq t\}$ be a Gaussian Markov process of orders $M=N=1$ and let $\{X(t), 0 \leq t\}$ be defined by

$$X(t) = H_p(\xi(t)) .$$

Then $E_{\mathbf{a}}(X(t)) = E(H_p(\xi(t) \mid \xi(u), u \leq s)) = H_p(E_{\mathbf{a}}(\xi(t))) = H_p(a(s, t)\xi(s)) = a^p H_p(\xi(s)) = a^p X(s)$.

Hence, $\{X(t)\}$ is the Markovian of non-linear order $M=1$. But $X(t) - a^P X(s)$ is orthogonal to $H_u(X)$: for all $u \leq s$, $E((X(t) - a^P X(s))X(u)) =$

$$EE(X(t)X(u)) - a^P E(X(s)X(u)) = a^P E(X(s)X(u)) - a^P E(X(s)X(u)) = 0.$$

So, $P_u(X(t)) = a^P X(s) = E_u(X(t))$ and $M=1$. \square

In the next example $E_u = P_u$, but $M \neq N$.

Example 3. We use some results from [3] and [8] in the following way:

Let the processes $\{Z_j(t), 0 \leq t\}$, $j=1, \dots, n$ be mutually orthogonal wide-sense martingals and let the function $\psi(t)$ satisfy some continuity conditions. Consider the second-order continuous process $\{\xi(t), 0 \leq t\}$ defined by

$$\xi(t) = \sum_{j=1}^n \psi^{j-1}(t) Z_j(t).$$

We have that $H_t(\xi) = \bigoplus_{j=1}^n H_t(Z_j)$, $0 \leq t$, and $P_u(\xi(t)) = \sum_{j=1}^n \psi^{j-1}(t) Z_j(s)$.

It follows that $\{\xi(t)\}$ is the Markovian of the linear order $N=M(s)=n$.

If $\{W(t), 0 \leq t\}$ is a Wiener process, then $\{H_k(W(t), 0 \leq t\}$, $k=1, \dots, n$ are mutually orthogonal wide-sense martingals. For the process $\{X(t), 0 \leq t\}$ defined by

$$X(t) = \sum_{j=1}^n \psi^{j-1}(t) H_j(W(t)),$$

we have

$$P_u(X(t)) = E_u(X(t)) = \sum_{j=1}^n \psi^{j-1}(t) H_j(W(s)).$$

So, the proces $\{X(t)\}$ is the Markovian of the linear order $N=n$ (H_u is the linear closure of $H_k(W(s))$, $k=1, \dots, n$) and of the non-linear order $M=1$ (ξ_u is generated by the random variables $W(s)$). \square

In the next example $M=N$, but $E_u \neq P_u$.

Example 4. Let $\{X(t), 0 \leq t\}$ be defined by

$$X(t) = \sum_{k=1}^m t^{\alpha-k} H_k(W(t)),$$

where $\alpha > m$. As

$$E_u(X(t)) = E(X(t) | W(u), u \leq s) = \sum_{k=1}^m t^{\alpha-k} H_k(E_u W(t)) = \sum_{k=1}^m t^{\alpha-k} H_k(W(s)),$$

the process $\{X(t)\}$ is the Markovian of the non-linear order $M=1$. The correlation function of $\{X(t)\}$ is, $s \leq t$,

$$r(s, t) = E X(s)X(t) = E \left[X(s) E_n X(t) \right] = E \left[\left[\sum_{k=1}^n s^{\alpha-k} H_k(W(s)) \right] \left[\sum_{j=1}^n t^{\alpha-j} H_j(W(s)) \right] \right] = \\ = \sum_{k=1}^n s^{\alpha-k} t^{\alpha-k} E H_k^2(W(s)).$$

Since $E H_k^2(W(s)) = c_k s^k$, c_k are constants, we have

$$r(s, t) = s^\alpha \sum_{k=1}^n c_k t^{\alpha-k} = f(s)h(t).$$

We conclude that $\{X(t)\}$ is the Markovian of the linear order $N=N(s)=1$.

It is easy to see that

$$P_n(X(t)) = \frac{h(t)}{h(s)} X(s) * E_n(X(t)). \quad \square$$

Using the last two examples, it is clear that the relation between M and N may be arbitrary.

Example 5. Let $\{W_j(t), 0 \leq t\}$, $j=1, \dots, m$, $m \leq n$ independent Wiener processes. The process $\{X(t), 0 \leq t\}$ defined by

$$X(t) = \sum_{k=1}^n t^{\alpha-k} H_k(W_{j_k}(t)), \quad j_k \in \{1, \dots, m\},$$

is the Markovian of $M=m$ and $N=1$.

The process $\{Y(t), 0 \leq t\}$ defined by

$$Y(t) = \sum_{k=1}^n \psi^{k-1}(t) H_k(W_{j_k}(t))$$

is the Markovian of $M=m$ and $N=n$. \square

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Rezime

NELINEARNI RED MARKOVSKOG PROCESA

Definišemo nelinearni red $M < \infty$ Markovskog procesa $\{X(t), 0 \leq t\}$ kao $M = \sup M(s)$ gde je $M(s)$ minimalni broj slučajnih promenljivih koje generišu $\sigma \{E_s X(t), s \leq t\}$ (E_s je uslovno očekivanje u odnosu na $\sigma \{X(u), u \leq s\}$). Diskutujemo odnos između nelinearnog i linearnog reda Markovskog procesa. Ova poslednji je jedna modifikacija Hida-nog višestrukog Markovskog procesa.

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